

II. THE ROLE OF ANTHROPOGENIC WARMING IN 2015 CENTRAL EUROPEAN HEAT WAVES

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Station-based observations and bias-corrected model simulations show that the frequency of short-term heat waves in central Europe has increased, albeit quantitative estimates of risk ratios differ considerably between methods.

Summer 2015 in Europe. The summer 2015 in Europe was highly unusual, as persistent heat and dryness prevailed in large parts of the continent. In central and eastern Europe, a combination of record-low seasonal rainfall (Orth et al. 2016) and record-high monthly July/August temperatures were observed over an area stretching from France to western Russia (Supplemental Fig. S11.1). The anomalous temperatures were caused by a sequence of four intense heat waves that struck the region from the end of June to early September (e.g., Fig. 11.1a). It is precisely the few-day heat that causes problems with human health, especially when combined with high humidity (McGregor et al. 2010). We analyze seasonal maxima of 3-day mean temperature ($T_{air,3d,max}$) and seasonal maxima of 3-day daily maximum wet bulb temperature ($WBTX_{3d,max}$), a measure of human thermal discomfort that combines temperature and humidity and is a proxy for heat stress on the human body (Fischer and Knutti 2013; Sherwood and Huber 2010).

The series of heat waves began with a strongly meandering jet stream, that is summertime “omega-blocking” (Dole et al. 2011), and the advection of very warm subtropical air into central and western Europe (Supplemental Fig. S11.1). Later in the season, the jet stream was displaced to the north, so that stable high-pressure systems could prevail over central and eastern Europe bringing heat there. The first heat wave in early July was hence most pronounced in western parts of the continent, while south-central and east-central Europe experienced the highest

temperatures in the subsequent heat waves later in the season (Fig. 11.1b).

Anomalies in the hottest 3-day mean temperature reached up to $+6^{\circ}\text{C}$ relative to climatology (Figs. 11.1c,d), and temperature records were broken, including nationwide records (Kitzingen, Germany: 40.3°C ; <https://weather.com/news/climate/news/europe-heat-wave-poland-germany-czech-august-2015>), various station records stretching from France to the Balkan countries and southern Sweden (www.meteofrance.fr/actualites/26913226-episode-de-tres-fortes-chaleurs-en-france), nighttime temperatures (Vienna, Austria: 26.9°C), record 3-day mean temperatures across central Europe (Fig. 11.1e), and inland water temperatures (e.g., Lake Constance). Europe experienced the hottest August ever recorded (NOAA 2016), and the entire summer season ranked third after the unusual summers of persistent heat in 2003 and 2010 with hotspots in France and western Russia, respectively (Barriopedro et al. 2011; Stott et al. 2004). This extraordinary sequence of events raises the question to what extent human-induced climate change played a role in short-term heat waves beyond natural climate variability.

A potential anthropogenic contribution to the summer 2015 heat events had already been investigated in near-real time (www.climatecentral.org/europe-2015-heatwave-climate-change), and in the present paper we build upon and substantiate the previous analysis. We investigate two diagnostics ($T_{air,3d,max}$ and $WBTX_{3d,max}$) at four locations in long-term station-based observational records and in a large ensemble of consistently bias-corrected regional climate model simulations.

Methods and Data. First, we analyze long-term observational data (115 years of data for each station) from the ECA&D dataset (Klein Tank et al. 2002) of four central and eastern European stations that

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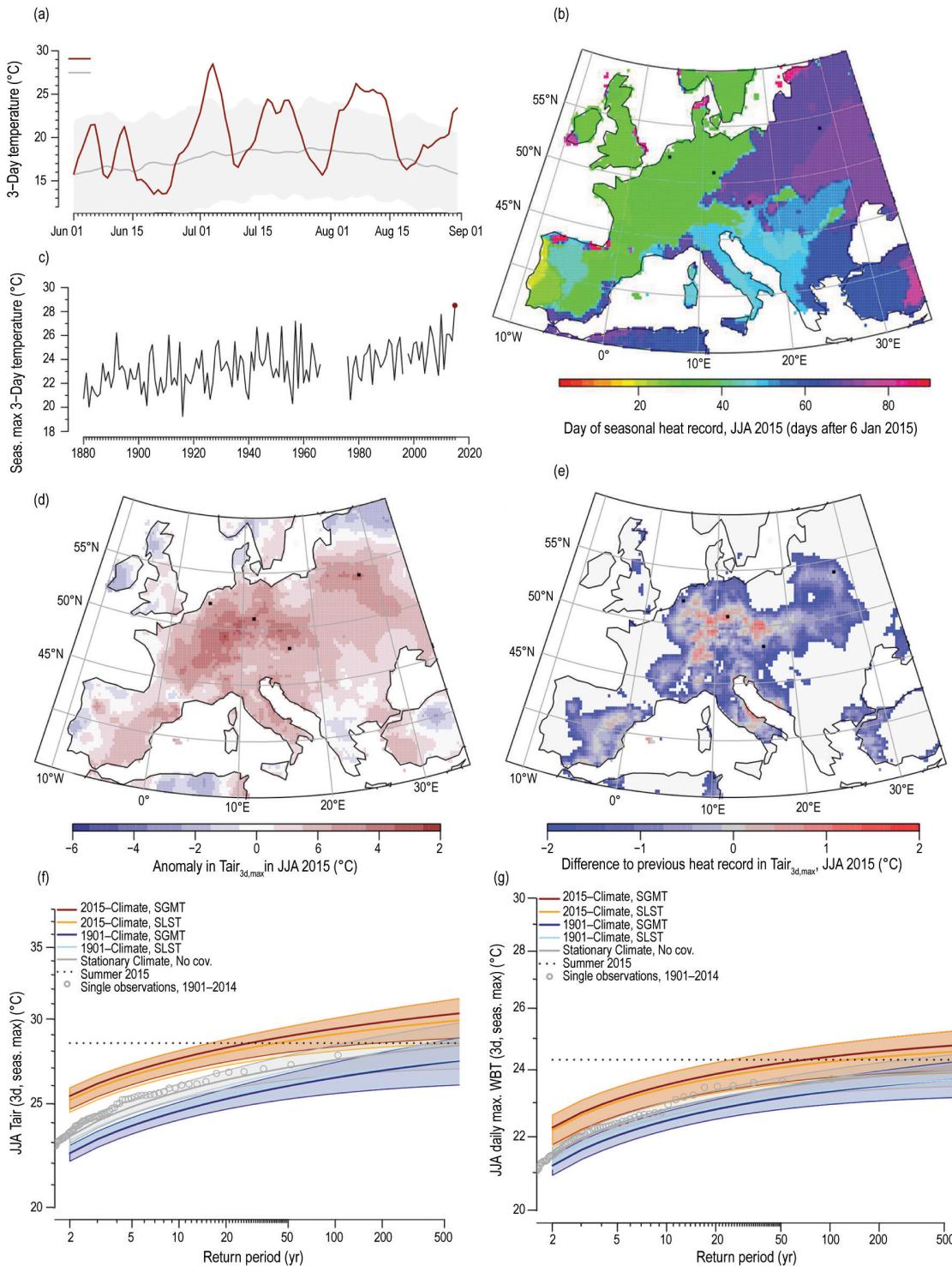


FIG. II.1. (a) Time series of 3-daily mean temperatures in summer 2015 at the Jena site (gray shading denotes $\pm 2\sigma$ deviations relative to long-term interannual variability). (b) Day of seasonal temperature record in summer 2015. (c) Time series of seasonal maximum of 3-day mean temperatures ($T_{air,3d,max}$) at the Jena site (summer 2015 is marked by a red dot). (d) Anomalies in $T_{air,3d,max}$ over Europe in summer 2015 relative to 1981–2010. (e) Difference to previous heat records (1950–2014) in $T_{air,3d,max}$ in the EObs dataset. Positive differences indicate a new heat record in JJA 2015. (f),(g) Return time plots of GEV fits for $T_{air,3d,max}$ and $WBTX_{3d,max}$, respectively, at the Jena site. Red (orange) lines indicate the fit for 2015 climate, dark-blue (light-blue) lines indicate the fit for 1901 climate for a smoothed global mean temperature covariate (smoothed local summer temperature covariate).

were affected by the heat waves in summer 2015 (Table 11.1), using data from 1901 onward. For each station, annual time series of $T_{air,3d,max}$ and $WBTX_{3d,max}$ are calculated for July–August. $WBTX_{3d,max}$ is derived from daily maximum air temperature and vapor pressure (computed from relative humidity and daily mean temperature; www.srh.noaa.gov/epz/?n=wxcalc_rh) using an iterative procedure based on the psychrometric equation (Sullivan and Sanders 1974). Subsequently, generalized extreme value (GEV) statistical models are fitted to the data (Coles 2001) excluding the year 2015, using two different assumptions about changes in climate:

1) A “local” covariate to the location parameter of the GEV (21-year smoothed local summer temperatures, SLST) as a proxy for any changes to local climate;

2) A “global” covariate to the location parameter (21-year smoothed global mean temperatures, SGMT) as a proxy for anthropogenic influence on climate (van Oldenborgh et al. 2012).

To avoid overfitting the relatively low number of data points, no dependence in the scale or shape parameter is assumed. Probability ratios (PR) based on the GEV as a metric to quantify human-induced change in the odds of extreme events ($PR = p_{ANT}/p_{NAT}$; Fischer and Knutti 2015) were obtained by calculating the probability of an event as warm or warmer than

the observed 2015-event in a 2015-climate (p_{ANT}), and in 1901 as a proxy for preindustrial climate.

Second, a model ensemble-based assessment using the global general circulation model HadAM3P ($1.875^\circ \times 1.25^\circ \times 15$ -min resolution) and a dynamically downscaled regional variant (HadRM3P, $0.44^\circ \times 0.44^\circ \times 5$ -min resolution) is conducted to complement the empirical analysis (see Massey et al. 2015 for all details regarding the model setup). Initial condition ensembles are generated for an anthropogenic scenario (ANT, $n = 2286$), in which the model is driven by observed (2015) sea surface temperatures (SSTs) and anthropogenic forcings in atmosphere-only mode for 1 year at a time (starting 1 December; Massey et al. 2015); and a natural scenario (NAT, $n = 4414$) with all anthropogenic forcings (i.e., greenhouse gases, aerosols, halocarbons, and ozone) set to pre-industrial levels and 11 different estimates of natural SSTs (Schaller et al. 2014). For each of the four locations (centered over a $1^\circ \times 1^\circ$ grid cell), a resampling bias correction strategy based on an observational constraint is applied to the model ensemble (Sippel et al. 2016) because the raw model output is notoriously too hot and dry (Black et al. 2015; Massey et al. 2015), severely compromising attribution statements (Supplemental Fig. S11.2). The seasonal maximum 21-day average temperature from the E-OBS dataset (Haylock et al. 2008) is used as a resampling

Table 11.1. Location of meteorological stations and probability ratios estimated from observed and simulated data. Very large PR with a lower bound (5% confidence interval) exceeding 10 are reported as >10. PR from the model output are given as 5th to 95th percentile of 100 bootstrapped replicates ($n = 1000$). A PR range exceeding one would be significant at 95% confidence under a one-sided test. PR for the original model simulations (i.e., non-bias corrected) are indicated for comparison only. *The observed De Bilt series contains a well-known inhomogeneity in 1950, so the homogenized series from KNMI was used instead. **Humidity data was not available for Vienna and Minsk in the ECA&D dataset for the year 2015.

Station	De Bilt*	Jena	Minsk	Vienna
Country	Netherlands	Germany	Belarus	Austria
Location	52°06'N, 5°11'E	50°55.5'N, 11°35'E	53°52'N, 27°32'E	48°14'N, 16°21'E
$T_{air,3d,max,2015}$ (°C)	25.2	28.5	27.3	29.1
$PR_{HadRM3P, BC-anom}$	1.2–1.4	1.1–2.5	1.7–2.5	1.8–2.9
$PR_{HadRM3P, BC-anom, obs. trend}$	4.7–7.5	4.1–8.7	3.4–5.2	>10
$PR_{EOBS, GEV-GMT}$	>10	>10	>10	>10
$WBT_{3d,max}$ (2015, °C)	22.9	24.3	n/a**	n/a**
$PR_{HadRM3P, BC-anom}$	1.3–1.8	1.5–3.1	n/a**	n/a**
$PR_{HadRM3P, BC-anom, obs. trend}$	>10	2.7–7.7	n/a**	n/a**
$PR_{EOBS, GEV-GMT}$	>10	>8.6	n/a**	n/a**

constraint, and a percentile-based transfer function is calibrated for each station separately on the 1986–2010 climatology using an identical model setup (Massey et al. 2015). Subsequently, both natural and anthropogenic simulations are resampled using the derived relationship (Sippel et al. 2016). In contrast to widely used methods like quantile–quantile mapping, resampling retains the full multivariate structure and physical consistency of the model output but reduces the available ensemble size and chooses colder and wetter ensemble members, therefore alleviating the hot and dry bias (Sippel et al. 2016). In the context of event attribution, it is applied for the first time in this paper (Figs. 11.2a–d; see next section). To avoid potential mean biases due to station location, the mean of the resampled ensemble is adjusted to the station mean (Supplemental Figs. S11.2c,d). Results are demonstrated exemplarily for one station (Jena), and probability ratios are reported for all stations.

Results and Discussion.

The statistical analysis of estimated return times of $Tair_{3d, max}$ reveals that 2015-like heat events occur in present day climate approximately every 27 years in Jena with the one-sided 5% lower confidence bound at 16 years (Fig. 11.1). Including both the local and global climate change covariates into the GEV fit demonstrates a profound increase in return times of those types of events relative to earlier years for both $Tair_{3d, max}$ and $WBTX_{3d, max}$ in Jena (Figs. 11.1f,g) and all other locations with probability ratios typically exceeding a value of 10 (Table 11.1). The intensity of heat waves increases by about 3°C in $Tair_{3d, max}$ but only 1.1°C in $WBTX_{3d, max}$ (Figs. 11.1f,g). In spite of this difference, the increase in the probability ratio is similar.

A similar analysis is conducted in a very large ensemble of model simula-

tions. The 21-day resampling constraint considerably improves the representation of short-term heat waves by avoiding physically implausible simulations (Figs. 11.2a–d) and improving the simulated variability of heat waves (Supplemental Figs. S11.2c,d). The correlation structure between the temperature constraint and short-term heat stress ($WBTX_{3d, max}$) in the observations is reproduced in the resampled model ensemble but not in the original model ensemble (Figs. 11.2a,c). This indicates that robust attribution statements for impact-related, and thus multivariate quantities (such as $WBTX_{3d, max}$), require a physically consistent bias correction of model output.

Consistent with the observations, the model-based assessment shows a shift in the return periods toward more frequent and more pronounced summer heat stress (Fig. 11.2b) in all locations (Table 11.1) and both bias-corrected and original simulations. The probability ratios derived from the bias-corrected

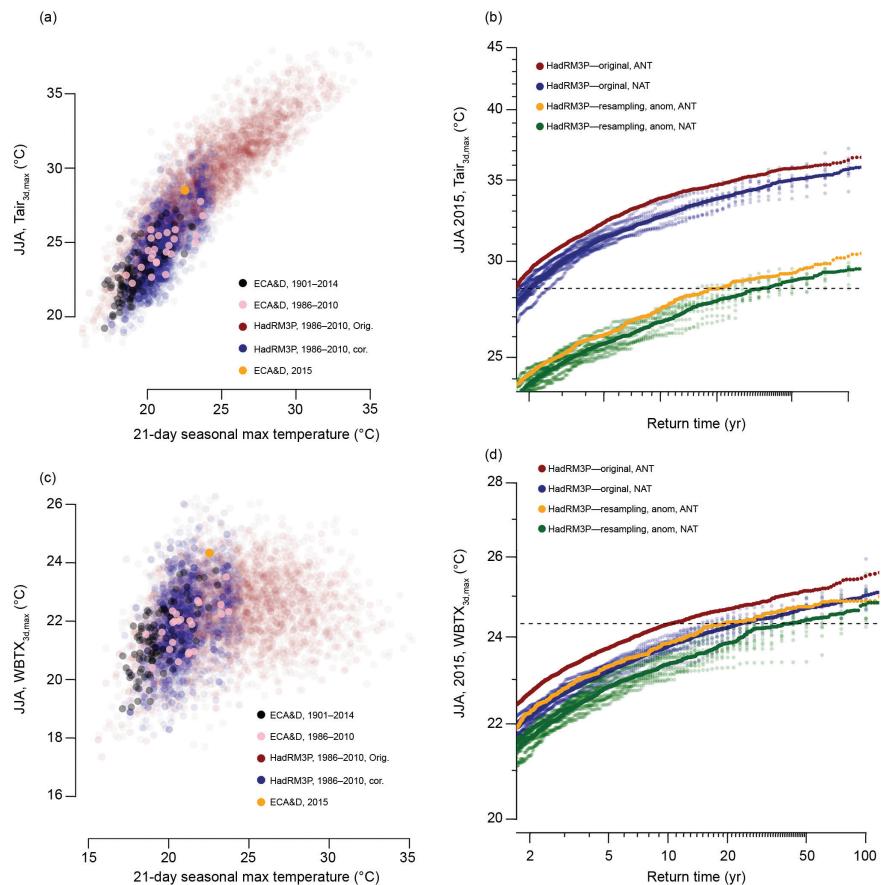


FIG. 11.2. (a),(c) Correlation between 21-day seasonal maximum temperature (observational constraint for resampling bias correction) and impact-related quantities ($Tair_{3d, max}$ and $WBTX_{3d, max}$, respectively). Pink dots correspond to 1986–2010, the period used for calibration of the bias correction resampling function. (b),(d) Return time plots for original and bias-corrected model output for $Tair_{3d, max}$ and $WBTX_{3d, max}$, respectively.

model ensembles range from 1.1 to 2.9 ($Tair_{3d, max}$) for the four locations ($PR = 1.3 - 3.1$ for $WBTX_{3d, max}$ in Jena and De Bilt), depending on the magnitude of the 2015 event, the model-simulated warming, and interannual variability. These estimates are thus lower than those estimated from the observations but can be largely explained by method- and data-related differences. For instance, the statistical method assumes that the trend is caused fully by anthropogenic factors, while the model analysis is based on a “real counterfactual” scenario but tends to underestimate warming trends in temperature extremes in Europe (Min et al. 2013). The mean observed change across all locations between 2015 and 1901 of $3.1^{\circ}C$ ($Tair_{3d, max}$) and $2.2^{\circ}C$ ($WBTX_{3d, max}$) is much larger than the original ($+1.1^{\circ}C$ in $Tair_{3d, max}$ and $+0.5^{\circ}C$ in $WBTX_{3d, max}$) and bias corrected ($+0.9^{\circ}C$ in $Tair_{3d, max}$ and $+0.5^{\circ}C$ in $WBTX_{3d, max}$) model simulations. Hence, replacing the model-simulated warming by the observed change between 1901 and 2015 causes roughly a tripling of probability ratios for the bias-corrected simulations at all locations (e.g., 3.4–8.7 for $Tair_{3d, max}$ and 2.7 to exceeding 10 for $WBTX_{3d, max}$; cf., Table 11.1). Furthermore, uncertainties due to event selection (Christiansen 2015), dependence on the spatial and temporal scale (Angélil et al. 2014), high nonlinearity in attribution metrics such as the probability ratio (Supplemental Fig. S11.2), and a slightly higher variability on sub-monthly time scales in the model simulations than in the observations despite bias correction further contribute to model-data discrepancies and variability in the presented estimates of the probability ratios.

Conclusion. In conclusion, the multimethod analysis applied in this paper provides consistent evidence that human-induced climate change has contributed to the increase in the frequency and intensity of short-term heat waves and heat stress such as the central and eastern Europe 2015 event.

However, quantitative estimates of the risk ratio at local scales can differ widely depending on the exact methodologies applied, thus highlighting large method- and data-related uncertainties. In this study, due to the large discrepancy between observed and modeled trends in temperature extremes, the model-estimated probability ratios are lower than those estimated from the observations.

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